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## V.I.Pagurova, N.S.Chizhikova (Moscow, Lomonosov MSU). On the joint limiting distribution of central and intermediate order statistics.

We consider the joint asymptotic distribution of central and intermediate order statistics when a sample size tends to infinity.

Let  $X_1, \ldots, X_n$  be mutually independent random variables with the common distribution function F(x), f(x) = F'(x), the set of order statistics is  $X_1^{(n)} \leq \cdots \leq X_n^{(n)}$ . Let  $0 < t_1 < \cdots < t_m, t_{m+2} > \cdots > t_{m+l+1} > 0$ ,  $0 < \alpha < 1$ , 0 , <math>[x] denotes the integer part of  $x, n_i = [t_i n^{\alpha}], i = 1, \ldots, m, n_{m+1} = [np] + 1$ ,  $F(\zeta) = p, n_j = [n - t_j n^{\alpha} + 1], j = m + 2, \ldots, m + l + 1$ . Define

$$\lambda_{k,n} = k/n, \quad k = k(n) \to \infty, \quad \lambda_{k,n} \to 0, \text{ as } n \to \infty.$$
 (1)

Necessary and sufficient conditions under which the statistic  $T_n = (X_k^{(n)} - d_n)/c_n$  has a Gaussian distribution, as  $n \to \infty$ , for some  $c_n > 0$  and  $d_n$  were shown in [1, 2]. If the condition (1) is satisfied then the the Gaussian and log-Gaussian distributions are possible limiting distributions [3, 4]. The joint asymptotic distribution, as  $n \to \infty$ , for central order statistics of a rank  $[n\lambda_i] + 1$ ,  $i = 1, \ldots, m, 0 < \lambda_1 < \cdots < \lambda_m < 1$ , was given in [5]. The joint asymptotic distribution of intermediate order statistics when random variables  $X_1, \ldots, X_n$  satisfy some conditions of dependence was shown in [6]. The asymptotic limit distribution of intermediate order statistics based on the sample with a random size was considered in [7].

We introduce variables  $d_{n,i}$  and  $b_{n,j}$  satisfying the following equations

$$F(d_{n,i}) = t_i/n^{1-\alpha}, \quad i = 1, \dots, m, \quad F(b_{n,j}) = 1 - t_j/n^{1-\alpha}, \quad j = m+2, \dots, m+l+1.$$

**Theorem.** Let f(x) be a differentiable in the neighborhood of  $d_{n,i}$ ,  $\zeta$ ,  $b_{n,j}$ ,  $f(d_{n,i}) \neq 0$ ,  $i = 1, \ldots, m$ ,  $f(b_{n,j}) \neq 0$ ,  $j = m + 2, \ldots, m + l + 1$ , and  $\lim_{n\to\infty} n^{1-\alpha/2} f(d_{n,1}) \neq 0$ ,  $\lim_{n\to\infty} n^{1-\alpha/2} f(b_{n,m+2}) \neq 0$ . Then for every  $\zeta$ ,  $f(\zeta) \neq 0$ , the joint distribution of random variables

$$X_{n_1}^{(n)} - d_{n,1}, \dots, X_{n_m}^{(n)} - d_{n,m}, X_{n_{m+1}}^{(n)} - \zeta,$$
  
$$X_{n_{m+2}}^{(n)} - b_{n,m+2}, \dots, X_{n_{m+l+1}}^{(n)} - b_{n,m+l+1}, \quad (2)$$

as  $n \to \infty$ , converges to (m+l+1)-variate Gaussian distribution with expectations equal

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zero and covariances in the following asymptotic presentations

$$\begin{aligned} & \cos\left(X_{n_{i}}^{(n)}, X_{n_{j}}^{(n)}\right) = t_{i}/(n^{2-\alpha}f(d_{n,i})f(d_{n,j})) \quad i, j = 1, \dots, m, \, i \leq j, \\ & \cos\left(X_{n_{i}}^{(n)}, X_{n_{j}}^{(n)}\right) = t_{i}/(n^{2-\alpha}f(b_{n,i})f(b_{n,j})) \quad i, j = m+2, \dots, m+l+1, \, j \leq i, \\ & \cos\left(X_{n_{i}}^{(n)}, X_{n_{m+1}}^{(n)}\right) = t_{i}(1-p)/(n^{2-\alpha}f(d_{n,i})f(\zeta)), \quad i = 1, \dots, m, \\ & \cos\left(X_{n_{j}}^{(n)}, X_{n_{m+1}}^{(n)}\right) = t_{j}p/(n^{2-\alpha}f(b_{n,j})f(\zeta)), \quad j = m+2, \dots, m+l+1, \\ & \cos\left(X_{n_{i}}^{(n)}, X_{n_{j}}^{(n)}\right) = t_{i}t_{j}/(n^{3-2\alpha}f(d_{n,i})f(b_{n,j})), \quad i = 1, \dots, m, \, j = m+2, \dots, m+l+1, \\ & \mathbf{D}X_{n_{m+1}}^{(n)} = p(1-p)/(nf^{2}(\zeta)). \end{aligned}$$

If covariances tend to zero then random variables (2) are independent asymptotically.

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